Artificial Intelligence formulated this projection for compatibility purposes from the original article published at Global Journals. However, this technology is currently in beta. *Therefore, kindly ignore odd layouts, missed formulae, text, tables, or figures.* 

# Greek Crisis, Stock Market Volatility and Exchange Rates in the European Monetary Union: A Var-Garch-Copula Model

 3
 Jaghoubbi Salma<sup>1</sup>

 4
 <sup>1</sup> University of Tunis el Manar

 5
 Received: 7 December 2013 Accepted: 3 January 2014 Published: 15 January 2014

## 7 Abstract

The main objectives of this study are twofold. The first objective is to examine the volatility 8 spillover between seventeen European stock market returns and exchange rate, over the period 9 2007-2011, in a multivariate setting, using the VAR (1)-GARCH (1,1) model which allows for 10 transmission in returns and volatility. The second is to investigate the dependence structure 11 and to test the degree of the dependence between financial returns using two measures of 12 dependence: correlations and copula functions. Five candidates, the Gaussian, the Student?s 13 t, the Frank, the Clayton and the Gumbel copulas, are compared. Our empirical results for 14 the first objective suggest that past own volatilities matter more than past shocks (news) and 15 there are moderate cross market volatility transmission and shocks between the markets. 16 Moreover, the result on the second objective implies that, considering all the financial returns 17 together, the Student-t copula seems the best fitting model, followed by the Normal copula, 18 both for the two sub-period. The dependence structure is symmetric and has non-zero tail 19 dependence. However, if we examine the relationship between each pair of stock-FX returns, 20 both of the degree of the dependence and the dependence structure vary when the financial 21 Greek crisis occurs. Our findings have important implications for global investment risk 22 management by taking into account joint tail risk. 23

24

Index terms— greek financial crisis, return spillover, volatility spillover; foreign exchange rate, var-garch (1,
 1)-copula model.

## 27 1 Introduction

nderstanding the dependence structure across international financial markets remains a crucial issue for risk 28 management and portfolio management. Several studies have focused on the comovement of world exchange 29 indices during a worldwide financial crisis. Moreover, many researchers have investigated the relationship among 30 worldwide financial markets. There is a great deal of research focusing on the co-movements of international 31 equity markets. Following the stock market crash of October 1987 in the United States, ??ing and Wadhwani 32 (1990) tried to explain why, in October 1987, almost all financial markets collapsed together despite different 33 34 economic contexts. In 1996, ?? alvo and Reinhart estimated that the co-movements of weekly equities returns and 35 Brady bonds, in Asia and Latin America, were higher after the crisis. ??aig and Goldfajn (1999) investigated the 36 links between five financial markets which are Thailand, Malaysia, Indonesia, Korea and the Philippines. They tested the statistical significance of the increase in correlation coefficients of exchange markets equity, interest 37 rate and sovereign debt. They confirmed the contagion effect only in Thailand and Malaysia. However; they 38 found that Thailand had not played an important role in the process of contagion during the Asian crisis. Forbes 39 and Rigobon (2002) attempted to test the existence of contagion effect during the following crisis: the U.S stock 40 market crash of 1987, the Mexican peso crisis in 1994 and the crisis in South East Asia in 1997 using daily return 41 data. They showed that the correlation between different countries is not significantly higher during crises. 42

Besides, other examples of research on the co-movements of equity markets can be found in Karolyi and Stulz 43 (1996) and ??ongin and Solnik (2001), while the methodology used is along the line of correlations and conditional 44 correlations. However, several empirical studies, such as Boyer and al. (1999), ??orbes and Rigobon (2001) and 45 ??orsetti and al. (2002) showed that the use of the high frequency financial series indicates three types of the bias, 46 because of heteroskedasticity, endogeneity and other omitted variables. Since these limitations of correlation-47 based models, research has started to use copulas to directly model the dependence structure across financial 48 markets. Works along this line include Rochand Alegre (2005) who tested different structures of dependence, 49 including different type of copulas: the Gaussian, the t-Student and seven other Archimedean copulas to model 50 the dependence of Spanish market returns. Their results reject the Gaussian copula in almost all cases and among 51 the nine structures considered. Moreover, the Student-t-copula provides the best results. Jondeau and Rockinger 52 (2006), Bartram and al. (2007) and Dimitris Kenourgios Aristeidis Samitas (2011) estimate the conditional 53 copulas in order to model the dependence between the major market indices. They report asymmetric extreme 54 dependence between equity returns. Boubaker, A., and Jaghoubi, S., (2011) employ the student-t-copula to model 55 the dependence structure of among a sample of U eight emerging and eight developed markets. Their results show 56 that this new approach proves more appropriate to describe the non-linear and complex dynamics of the financial 57 58 market returns than traditional modeling which imply a normality hypothesis. In addition, they confirm the 59 contagious nature of the Subprime crisis between emerging and developed markets. While the above literature 60 focuses on the dependence structure and co-movements in equity markets via copulas, Okimoto (2008) also 61 employs copulas to model the asymmetric exchange rate dependence between US-UK and find that this regime is best described by asymmetric copula with lower tail dependence. Although there is wide literature analyzing the 62 co-movements and the interdependence between the international equity markets and some literature on modeling 63 the dependence structure between the exchange rates via copulas, few use copulas to study the co-movements 64 across markets of different asset types, such as the stock market and foreign exchange rates. 65 The purpose of this paper is to examine the dynamic correlation and volatility transmission between the 66

European Monetary Union and the FX returns and to explore the dependence structure between daily stock 67 returns, after the occurrence of the current financial Greek crisis. Our paper has similarities and differences 68 with the previous literature. The main similarity is that we try to estimate dependence of financial markets. 69 However, there are several main differences. First of all, while previous empirical investigations of the link 70 between FX markets and stock prices are mainly devoted to developed markets, and sometimes to Pacific Basin 71 72 countries, our interest is focused on European markets that are member of euro-area and were affected by 73 recent financial Greek crisis. Second, we assess dependence using both correlation and copula functions, and we are agnostic ex ante about which technique is appropriate. Third, unlike most studies in the literature 74 that directly model the dependence structure between FX rates and stock prices, using copula approach, we 75 attempt to estimate this dependence by combining two models which are the VAR-GARCH(p,q) and the Copula 76 techniques to have a joint VAR-GARCH-COPULA model with possibly skewed, fat tailed return innovations and 77 non-linear property. Although, the vector autoregressive generalized autoregressive conditional heteroskedasticity 78 model (VAR-GARCH) is used to explore the joint evolution of conditional returns, volatility and correlation 79 between the European stock market returns and the exchange rate over the Greek crisis period, the multivariate 80 dependence structure between markets is modeled by several copulas which are perfectly suitable for non-normal 81 distributions and nonlinear dependencies. 82

The remainder of this paper is organized as follows: Section 2 presents the theoretical background of the dependence measures used in empirical finance and shows how they can be applied to study the extreme comovements between the European markets. In Section 3, the empirical results are reported and interpreted. We provide summary of our conclusions in Section 4.

87 II.

## $^{88}$ 2 Methodology

Our methodology is based, primarily, on the calculation of linear and rank correlation coefficients between the 89 European market returns. We get series of correlation coefficients between these markets and we study their 90 91 dynamics changes. Secondly, such as measurements based on linear correlation may lead to misspecification of the dependence structure with its nonlinear portion, copula approach is employed to provide the robust measures 92 of dependences based on the entire joint distributions of variables and also to estimate dependence focuses on the 93 entire structure rather than correlation. Besides, as the copula functions are used to separate the margins and 94 the dependence structure corresponding to a joint distribution, we estimate, in the first step, the parameters of 95 marginal distributions and those of returns and volatilities equations. Then, in the second step, the parameters 96 of the copula taking into account the parameters estimated in the first step. 97

## <sup>98</sup> **3** a) Correlations

<sup>99</sup> Correlations are the most familiar measures of dependence in finance. Although most studies have focused <sup>100</sup> on measuring the dependence between financial markets have used the Pearson correlation, this coefficient is <sup>101</sup> only reliable when the random variables are jointly Gaussian. Therefore, we consider two other measures of <sup>102</sup> dependence: the Kendall's tau and the Spearman's Rho, which are measures of concordance, generalize the linear correlation, taking into account the joint distribution (and not just marginal) and are dependent on copulas. The rate of Kendall and Spearman's rho are two measures of concordance well known in statistics. They provide a measure of the correlation between the ranks of the observations, unlike the linear correlation coefficient which assesses the correlation between the values of observations. It is necessary to recall the notion of concordance. Let (x, y) and (?? ?, ?? ?) two realizations of a continuous random vector (X, Y), then (x, y) and(?? ?, ?? ?)are called concordant if (x-?? ?) (y-?? ?) > 0 and discordant if (x-?? ?) (y-?? ?) < 0.

? The Kendall correlation coefficient Let (X, Y) a couple of random vectors and (?? ?, ?? ?) a copy of (X, Y) that is to say a pair of vectors in all respects identical to (X, Y) the Kendall's tau is then:?? ?? (X, Y) = Pr111 {(?? ???)(?? ???) > 0} -Pr {(?? ???)(?? ???) < 0} (3.2)

<sup>112</sup> The Kendall's tau is simply the difference between the probability of concordance and of discordance.

? The Spearman correlation coefficient Let X and Y are two random variables of marginal distributions ?? ??
?????? ?? . The correlation coefficient Spearman rank coefficient ?? ?? is the Pearson correlation??between??
?? (??)and?? ?? (??) :?? ð ??"ð ??" (??, ??)= ?? (?? )??? (??)) (3.3)

b) A copula model for asymmetry dependence

<sup>117</sup> Copulas are multivariate distribution functions with standard uniform marginal distributions. Amdimensional <sup>118</sup> copula is represented as follows: C (u) = C (?? ?? , ?? ?? ) (3.5)

Where ?? 1, ??, ?? ?? are standard uniform marginal distributions. In such a context, copulas can be used to link margins into a multivariate distribution function. The copula function extends the concept of multivariate distribution for random variables which are defined over ??0,1]. This is possible due to the Sklar (1959) theorem which states that copulas may be constructed in conjunction with univariate distribution functions to build multivariate distribution functions. Sklar's Theorem: Let ?? ???? be a joint distribution function with margins ?? ?? and ?? ?? . Then there exists a copula C such that for all x, y in R,C (?? ?? ,?? ?? ) = C ( ?? ?? (x), ?? ?? (y)) = F (?? ?? ??? (?? ??),?? ?? ??? (?? ?? )) (3.6) C (?? ?? ,?? ?? ) = F (x, y)

126 If ?? ?? and ?? ?? are continuous, then C is unique; otherwise, C is uniquely determined on Ran ?? ?? ×Ran 127 ?? ?? and C is invariant under strictly increasing transformations of the random variables. Our model aims at 128 capturing the type of asymmetric dependence found in financial markets. For that, two models are specified: the 129 marginal distribution model and the joint distribution model.

## <sup>130</sup> 4 i. Specification of the marginal distribution

For marginal distributions, we use a bivariate VAR(1)-GARCH(1,1) 1 model developed by Ling and McAleer (2003)which allows for spillover effects in both returns and conditional volatilities to examine both own conditional volatility for each market and conditional cross market volatility transmission among European Monetary Union (EMU) and the FX rate. The conditional mean equation of the VAR (1)-GARCH (1, 1) system is giving by: ? ?? ?? = ?? + ??? ????? + ?? ?? ?? = ?? ?? ????? ?? ?? (3.7)

From these two equations above, we can see how volatility is transmitted over time across the EMU and the FX markets. Thus, the past shock and volatility of one market are allowed to impact the future volatility not only of itself but also of all other markets in the system.

ii. Specification of the dependence structure Here we study five copulas with different dependence structure: 139 the Gaussian copula, the Student-tcopula, the Frank copula, the Clayton and the Gumbel copula. From them, 140 the Gaussian copula is the most popular in finance and used as the benchmark. The following table shows 141 the characteristics of the best known models where the parameter C R ???? the distribution function of joint 142 variables, ?? ???? the degree of freedom, ? is the variance-covariance matrix, the parameter??measures the 143 degree of dependence between risks. ) = ? ?? ?? ?1 (?? 1 ), ? ? ? , ? ?1 (?? ?? )? Student ??, ?? ?? (?? 1 , 144 145 ??? ? 1) ? 1 ?? Gumbel ?? ? 1 C  $(u, v, ??) = \exp [?[(?Ln (u)) ?? + ??Ln (v)) ?? ? ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [?[(?Ln (u)) ?? + ??Ln (v)) ?? ? ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [?[(?Ln (u)) ?? + ??Ln (v)) ?? ? ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [?[(?Ln (u)) ?? + ??Ln (v)) ?? ? ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [?[(?Ln (u)) ?? + ??Ln (v)) ?? ? ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [?[(v, v, ??) + ??Ln (v)) ?? ? ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [?[(v, v, ??) + ??Ln (v)) ?? ? ] ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [?[(v, v, ??) + ??Ln (v)) ?? ? ] ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [?[(v, v, ??) + ??Ln (v)) ?? ? ] ] 1 ?? Frank ?? ? 0 C <math>(u, v, ??) = \exp [v, v, r, r]$ 146 v,?? = -1 ?? Ln [1+ (exp (??? ?? ) ?1 )( exp (??? ?? ) ?1 ) exp (??? )?1 III. 147

148 Data and Results

## <sup>149</sup> 5 a) Descriptive statistics

We use daily market data from seventeen European stock market indices, for a sample period of February 150 1, 2007 to December 21, 2011. We choose this period to investigate the impact of the 2009 Greek crisis on 151 the rest of European monetary countries. The countries used in our sample are France (CAC40), Germany 152 (DAX), Belgium (BEL-20), Spain (IGBM), Ireland (ISEQ), Italy (FTSE MIB), Luxembourg (LUX GENERAL), 153 the Netherlands (AEX), Ostrich (ATX), Portugal (PSI20), Finland (OMX H25), Greece (ATHEN COMPS), 154 155 Slovenia (SBI TOP), Cyprus (CYSE GENERAL), Malta (MSE), Slovakia (SAX) and Estonia (OMXT). The total number of observations is 1253 for the full sample. We briefly overview summary statistics, then discuss 156 the correlation and copula estimates. showing the probability of concordance is significantly higher than the 157 probability of discordance. The Spearman's Rhos for the pairs in each country are also positive for eleven 158 countries from seventeen which are France, Germany, Belgium, Ireland, Italy, Cyprus, Estonia, the Netherland, 159 Finland, Portugal and finally Luxemburg. However, the Spearman Rhos are negative for the rests of European 160 markets. From these results, we can conclude that there are strong rank correlations. 161

The German pair has the strongest dependence, followed by the Finland pair and the French pair. However, the 162 weakest is in the Spanish pair. In a e 4, we present these linear correlations and rank correlations measures for each 163 stockexchange rate return pair after the current financial Greek crisis. The linear correlation, Pearson coefficients, 164 for our pairs of returns are all positive, expect for the Cyprus market, showing that, for these sixteen European 165 markets, the increase (decrease) of the local stock market is associated with the appreciating (depreciating) of the 166 exchange rate EURO/USD. Besides, for the Cyprus, when the CYSE price increase (decrease), the EURO/USD 167 exchange rate depreciate (appreciate). Thus, the Cyprus stock market return and the exchange rate evolve in a 168 reverse sense. The Kendall's Taus for our pairs are all positive expecting for Cyprus, Luxemburg and Slovakia 169 indicating that the probability of concordance is higher than the probability of discordance. The Spearman's 170 Rhos indicate strong rank correlations. The values of Taus and Rhos are consistent with each other and the linear 171 correlation. The Spanish market has the strongest dependence with the EURO/USD exchange rate, followed by 172 the French pair, and the weakest is the Cyprus which has a negative dependence with the exchange rate. Further, 173 the correlation increase and became strong in the postcrisis period. Thus, the stock-exchange rate returns become 174 more dependent when financial extreme events (Greek crisis) occurs. 175

## <sup>176</sup> 6 c) Copula results

As the copula model allows us to separate the marginal behavior from the dependence structure, the estimation of copula models is decomposed into two steps: the first for the marginals and the second for the copulas. We employ the VAR-GARCH model for the marginal distributions of each stock index return and exchange rate return series. For the Joint model, we employ copulas with different dependence structure.

## <sup>181</sup> 7 i. Results of the marginal models

Our objective is to examine both own conditional volatility and shocks and conditional cross-market volatility 182 183 transmission and shocks between the Eurozone stock returns and the foreign exchange rate returns. For that, we use the Euro Stoxx 50 2 stock index for Eurozone (EMU) stocks and the EUR\_USD returns for the foreign 184 exchange market. We experiment on GARCH terms up to p=1 and q=1. The optimal lag order for the VAR 185 model is selected using the AIC and SIC information criteria. The estimation of the bivariate VAR (1)-GARCH 186 (1, 1) for the two sub-period, is presented in table 5. 2 We will discuss the empirical results of bivariate VAR(1)-187 GARCH (1,11) models in terms of own volatility and shock dependence, cross market volatility and shock spillover 188 for the Eurozone stock index and the FX rate, both for the pre-crisis and the post-crisis. During the pre-crisis 189 190 period and for the EMU, the sensitivity to past own conditional volatility and cross market volatility transmission are significant at the level of 1%, showing that future volatility can be predicted by both the past own conditional 191 volatility in the long run and the cross market volatility spillover. We found the same result for the own shocks 192 or news and cross market shock transmission, indicating a short run persistence. However, the effect of past 193 194 volatilities is much bigger than the effect of past shocks. This implies that fundamentals matter more than news. Considering now the FX rate, only the past own volatility is significant but has a negative coefficient, displaying 195 that own shocks and cross market volatility transmission and shocks cannot be used to predict either the future 196 volatility in the long run and the short run persistence. After the occurrence of the Greek crisis, the behavior 197 of these markets changes considerably. Indeed, the cross market volatility and shocks remains significant for the 198 EMU stocks but their persistence diverge. The results show the effect of past shocks on the Eurozone (EMU) 199 become bigger after the crisis, in contrast with the past own shocks effects', showing that news coming from the 200 FX market affect more returns dynamics than past own EMU news. Moreover, cross shocks (or spillover) are 201 more widespread inter-markets after the crisis. However, for the FX market, both own shocks and cross shocks 202 203 become significant at different level and have a positive effect in the short run. This finding show that past own shocks and shock spillover can be used in predicting future shocks or new. Besides, the foreign exchange market 204 becomes more sensitive to past shocks related to changes in news or noise than fundamentals. 205

ii. Results of the joint copula models We now present results from our copula estimation. We consider five
multivariate copulas, the multivariate normal, multivariate Student-t, multivariate Gumbel, multivariate Clayton
and the multivariate Frank. We first discuss the dependence structure using information criteria for European
stock markets and exchange rates. Table 6 report results from AIC, SIC and HQIC information criteria.

For the pre-crisis period, the best model which has lowest AIC, SIC and HQIC is the multivariate Student-t copula, with an average AIC of -8060.71, a SIC of -7450.90 and a HQIC of -7879.46 across countries, closely followed by the multivariate Gaussian copula. In the post-crisis period, the lowest AIC of -6780.04 corresponds to the Student-t copula, followed closely by the Gaussian model. The same results for the SIC and the HQIC information criteria. Thus, according to AIC, SIC and HQIC, the best fitting copula is the Student-t with symmetric tail dependence for the two sub-periods.

To better assess the degree as well as the dependence structure in the euro area, we will examine the relationship between each pair of stock-FX return separately, for the two sub period.

Table 7.A bellow, reports parameters estimates of bivariate copulas for each pair, before the occurrence of the financial Greek crisis. We note that the parameter?? ??and ?? measure the degree of dependence between returns and DoF is the degree of freedom in the Student-t copula. For all pairs, the dependence parameters; the correlation coefficient ? in both Gaussian and Student-t copulas, the degree of freedom DoF in the Student-t

copula and the asymmetric dependence parameter ?? in the Clayton, Gumbel and Frank copulas are positive 222 with the expect for Malta, Slovakia and Slovenia in the precrisis period. The correlation coefficient ? from the 223 Gaussian or Student-t copula is close to the usual correlation coefficient. The DoF of the Student-t copulas 224 225 are from 4 to 12, indicating the presence of extreme comovements and tail dependence. The tail dependence parameter ?? for pre-crisis period is 1.026 for the Luxemburg-foreign exchange rate pair. Thus, we can conclude 226 that only the LUX/EUR\_USD pair has asymmetric tail dependence. All the other stock market returns have 227 elliptical symmetric dependence structure (the case of the Gaussian or the Student-t copulas) with the foreign 228 exchange rate. 229

In order to appreciate both, the dependence structure and the degree of this dependence, after the Greek 230 crisis; we estimate the copula parameters in the post-crisis period. For all pairs, the dependence parameters; the 231 correlation coefficient ? in both Gaussian and Student-t copulas, the degree of freedom DoF in the Student-t 232 copula and the asymmetric dependence parameter ?? in the Clayton, Gumbel and Frank copulas are positive, 233 expect for Cyprus and Luxemburg. 234

The Spain return has the highest correlation coefficient with ? = 0.4277. The DoF of the Student-t copulas are 235 from 7 to 40, indicating the presence of strongly extreme co-movements and tail dependence. The tail dependence 236 parameter ?? for post crisis period, are from 0.02788 to 1.36. The French pair has the highest tail dependence 237 after the crisis, followed by the Ostrich pair and the Netherland pair. Moreover, the dependence structure 238 239 between each stock index returns and exchange rate returns is largely changed from a symmetric structure with 240 or not symmetric tail dependence to an asymmetric structure with non-zero and asymmetric upper and lower tail dependence. From our results, we find The Gumbel copula which is limited to the description of a positive 241 dependence structure. Thus, it allows only positive dependence structures or upper tail dependence, for which 242 the parameter belongs to the interval ??1,+?). We find also the Clayton copula which possesses similar properties 243 to the Gumbel copula. Consequently, the degree of the dependence varies when the financial Greek crisis occurs. 244 Indeed, as we see in tables above, it increased after the crisis, expect of Cyprus which remains symmetric but 245 with zero tail dependence. The degree of the dependence becomes weaker and moves from a positive to a negative 246 one. 247

Our findings may have important implications in the risk management. First, symmetric dependence structure 248 with zero tail dependence can specify different levels of correlation between the marginals; however, it must possess 249 radial symmetry which doesn't allow to extreme values correlation. Thus, in this case, the dependence has the 250 linear correlation coefficient as measure of dependence. Second, asymmetric dependence structure can have upper 251 252 tail dependence, lower tail dependence, or both; as such, they can better describe the reality of the behavior of financial markets. Additionally, it indicates the potential of simultaneous extreme events in both the stock 253 and foreign exchange market. This property of dependence structure is important to international investors who 254 invest in foreign stock markets. 255 IV.

### 256

#### Conclusion 8 257

This paper examines the dynamics relationship between foreign exchange and stock markets in the Economic 258 European Market, after the occurrence of the Greek crisis, using daily data from February 2007 to December 2011. 259 Based on the VAR(1)-GARCH(1,1) model, the results show that past own volatilities matter more than past 260

shocks (news) and there exist moderate cross market volatility transmission and shocks between the markets, 261 indicating that the past innovation in stock market have great effect on future volatility in foreign exchange 262 market and vice versa.

1

Noun Parameters Copulas Gaussian ??

?? ?? (?? 1 , ? ? . ?? ??

Figure 1: Table 1 :

 $\mathbf{2}$ 

Greek Copula Model

[Note: b] Empirical resultsi. Correlation estimates of dependence]

Figure 2: Table 2 :

## 8 CONCLUSION

## 3

Stock and FX	Mean	S.D	Skewness	kurtosis	Jarque-Bera
returns					
CAC40	-0.021238	0.779427	0.139538	7.969773	$1289.906 \ [0.000]$
DAX	-0.005244	0.746523	0.118501	8.201018	1414.069 [0.000]
<b>BEL-20</b>	250412.4	138032.6	-0.370712	2.493563	$2286.917 \ [0.000]$
IGBM	-0.022659	0.785953	0.240945	8.821721	$1777.869 \ [0.000]$
ISEQ-20	-0.041144	0.905208	-0.400758	8.079340	$1379.045 \ [0.000]$
FTSE-	-0.038663	0.819882	0.042126	7.258653	$944.1836\ [0.000]$
MIB					
LUXx	0.001546	0.000965	1.588051	4.654095	$649.3023 \ [0.000]$
ATX	1.152728	24.24244	20.33048	415.0361	8928443 [0.000]
PSI 20	-0.027363	0.643563	-0.013060	9.880040	$2465.221 \ [0.000]$
OMX	-0.016295	0.793113	0.101068	5.686303	$377.7563 \ [0.000]$
H25					
ATHEN.	-0.068151	0.901909	0.168730	6.378040	$602.4575 \ [0.000]$
COM-					
POS					
SBI-	0.046487	0.378522	-0.170289	10.76016	$3145.580 \ [0.000]$
TOP					
CYSE	0.018518	0.974732	-0.021023	7.124780	885.2031 [0.000]
MSE	-0.017356	0.305635	0.065548	9.336490	$2088.362 \ [0.000]$
SAX	0.023580	0.595649	1.592428	42.12614	$60751.30 \ [0.000]$
OMXT	-0.022058	0.657041	0.165734	8.714507	$1708.529 \ [0.000]$
EURO/USD.05E-05		0.316340	-0.192636	6.300257	577.9667 [0.000]
		Ostrich, Athens,	Spain, Sloven	ia, Portugal a	nd Slovakia;

Figure 3: Table 3

Pairs	Pearson correla- tion	Kendall's Tau	Spearman's Rho
French pair	$0.217002^{*}$	$0.120057^{*}$	$0.178925^{*}$
German pair	$0.241257^{*}$	$0.127205^{*}$	$0.191042^{*}$
Maltin pair	-0.013309	-0.002404	-9.58?? ?0.4
Belgium pair	$0.159411^*$	0.08665	$0.128747^{*}$
Irish pair	0.015392	0.021143	0.034781
Austrian pair	-0.053048	-0.050691	$-0.072301^{***}$
Greek pair	-0.038951	-0.022945	-0.029832
Italian pair	$0.093629^*$	0.048453	$0.075733^{**}$
Spanish pair	-0.08543**	-0.052023	-0.073413
Slovenian pair	0.007018	-0.002555	-0.001342
Cyprus' pair	0.048925	0.005701	0.0095
Estonian pair	$0.118461^*$	0.059192	$0.090004^*$
The Netherland's pair	$0.214112^{*}$	$0.120367^*$	$0.179963^*$
The Finnish pair	$0.222949^*$	$0.148791^*$	$0.219887^*$
Luxemburg's pair	0.035408	0.021129	0.037361
Portugal's pair	0.020022	-8.26 ?? ?0.4	0.02933
Slovaquie pair	0.014043	-0.005785	-0.007046

[Note: This table gives different correlation measures for each stock-EUR/USD exchange rate daily return pair over the period February 1, 2007 to October 15, 2009. \*, \*\*, \*\*\* denote significance level at the 1%, 5% and 10% respectively. Total observations are 691.]

Figure 4: Table 3 :

## $\mathbf{4}$

Copula Model 55

Figure 5: Table 4 :

Copula Model Year 2 ( )

Figure 6:

 $\mathbf{5}$ 

Figure 7: Table 5 :

## 3

## 6

Models	SIC	AIC	HQIC
Panel A: Pre-crisis			
Clayton	-1266.40	-1270.93	-1269.18
Gumbel	-1093.80	-1098.33	-1096.58
Normal	-6894.37	-7500.95	-7320.14
Student-t	-7450.90	-8060.71	-7879.46
Frank	-972.22	-976.75	-975.00
Panel B: Post-crisis	SIC	AIC	HQIC
Clayton	-1144.87	-1149.20	-1147.51
Frank	-888.76	-893.08	-891.40
Gumbel	-1001.04	-1005.36	-1003.68
Normal	-6033.35	-6580.57	-6437.33
Student-t	-6230.29	-6780.04	-6636.91

Figure 8: Table 6 :

## 7

Pairs	Copula models	?	Parar DoF	net SHC	Information criteria AIC	HQIC
France	Student-t	0.1871		-32.08	-41.14	-37.65
German	Student-t	0.1997	6	-34.16	-43.22	-39.72
Ostrich	Student-t	0.2388	5	-51.96	-61.02	-57.52
Belgium	Student-t	0.1347	5	-24.18	-33.24	-29.75
Netherland	Student-t	0.1882	5	-31.60	-40.65	-37.16
Athens	Student-t	0.2102	4	-47.92	-56.98	-53.48
Malta	Gaussian	-0.001325		11.17	2.11	5.60
Slovakia	Student-t	-0.005494	12	9.37	0.31	3.80
Cyprus	Gaussian	0.01002		6.18	1.65	3.40
Spain	Student-t	0.1805	5	-33.39	-42.45	-38.95
Ireland	Student-t	0.1085	4	-33.60	-42.66	-39.17
Luxemburg	Gumbel			1.0 <b>26</b> .28	4.22	7.71
Italy	Student-t	0.07914	4	-20.55	-29.61	-26.12
Finland	Student-t	0.2297	5	-47.34	56.39	52.90
Estonie	Student-t	0.09449	6	-7.96	-17.02	-13.53
Portugal	Student-t	0.003059	7	3.68	-5.38	-1.88
Slovenia	Gaussian	-0.001467		16.26	7.20	10.69

Figure 9: Table 7 .

	Copula Model Table 7.B : Estimation of copula parameters for the post-crisis period							
YeaPairs France		* *	eters for ?	Parameters DoF	210d ?	Information		
		Gumbel	-		1.36	criteria SIC		
						AIC HQIC		
						-85.73 -94.37		
						-91.01		
2	German	Student-t	0.3699	7		-73.82	-82.46 -79.10	
	Ostrich	Gumbel			1.331	-80.20	-88.84 -85.48	
	Belgium	Student-t	0.353	7		-70.48	-79.12 - 75.76	
	Netherland Gumbel				1.317	-67.56	-76.20 - 72.84	
	Athens	Gaussian	0.3599			-69.99	-74.32 - 72.63	
	Malta	Student-t 0.0066	534	40		13.06	4.42	
	SlovakiaClaytonCyprusGaussian -0.07477			0.0278	383.24	4.60		
				2.62				
	Spain	Gaussian	0.4277			-98.60 -102.92	-	
	Ireland	Clayton			0.471	-38.85	-47.49 -44.13	
	Luxemburg Gaussian -					7.21	2.89	
	Italy	Gumbel			1.139	-20.59	-29.23 -25.87	
(	Finland Estonie	Student-t	0.3576			-68.87 -6.05	-77.51 -74.15	
)		Gaussian	0.0101	1			1.73 - 3.41	
	Portugal	Clayton			0.2614	4-20.45	-29.09 - 25.73	
	Slovenia	Gaussian	0.0672	9		3.57	-0.76	

Figure 10:

## 8 CONCLUSION

 $<sup>^{1}</sup>$ © 2014 Global Journals Inc. (US)  $^{2}$ © 2014 Global Journals Inc. (US) tabl

- [Ling and Mcaleer ()] 'Asymptotic theory for a vector ARMA-GARCH model'. S Ling , M Mcaleer . *Econometric Theory* 2003. 19 p. .
- [Baig and Etgoldfajn (1999)] T Baig , I Etgoldfajn . Financial market contagion in the Asian Crisis, 1999. June.
   46.
- <sup>267</sup> [Calvo and Reinhart ()] 'Capital flows to Latin America: is there of contagion effects'. G Calvo , C Reinhart .
   <sup>268</sup> Private capital flows to Emerging Markets After the Mexican Crisis, (Washington, D.C) 1996. / Institute for
   <sup>269</sup> International Economics
- [Boubaker and Jaghoubi ()] Detecting financial markets contagion using copula functions, A Boubaker , S
   Jaghoubi . 2011. p. .
- [Putnam, S. (ed.)] Economic Interdependence and Flexible Exchange Rates, Putnam, S. (ed.) Cambridge, MA:
   MIT Press.
- [Login and Solnik ()] 'Extreme correlations of international equity markets'. F Login , B Solnik . The journal of
   *Finance* 2001. 56 p. .
- [Dimitris ()] 'Financial crises and stock market contagion in a Multivariate Timevarying Asymmetric Framework'.
   K Dimitris , AristedisS . Journal of International Financial 2011. 21 p. .
- 278 [Kenen, P.B. (ed.)] Handbook of International Economics, 2. Amsterdam, Kenen, P.B. (ed.) Elsevier.
- [Karolyi and Stulz ()] G A Karolyi , R M Stulz . Why Do Markets Move Together? An Investigation of the,
   (U.S.-Japan Stock Return Comovements) 1996. 51 p. .
- [King and Etwadhwani ()] M King , S Etwadhwani . Transmission of volatility between stock markets, 1990. 3 p.
   .
- [Chan et al. ()] 'Modelling multivariate international tourism demand and volatility'. F Chan, C Lim, M Mcaleer
   *Tourism Management* 2005. 26 p. .
- [Forbes and Rigobon ()] 'No contagion, only interdependence: Measuring stock market co-movements'. K J
   Forbes , R Rigobon . Journal of Finance 2002. 57 p. .
- [Roch and Alegre ()] 'Testing the bivariate distribution of daily equity returns using copulas. An application to
   the Spanish stock market'. O Roch , A Alegre . Computational Statistics & Data Analysis 2006. 51 p. .
- [Bartram et al. ()] 'The Euro and European financial market dependence'. S M Bartram , S J Taylor , Y H Wang
   Journal of Banking and Finance 2007. 31 p. .