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Haw the Default Probability is Defined by The Credit Portfolio Models: A Comparative Theoretical Analysis between the Structural Models?

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Abstract - This paper is elaborate of which the main is to present a theoretical analysis between the structural models. There are currently three types of models to consider the risk of credit: the structural models (The KMV Moody's model and the CreditMetrics model) also defined by the models of the value of the firm, reduced form models also defined by models with intensity (actuarial models) and the econometric models (The macro-factors models). The development of its three types of models is based on a theoretical basis developed by several researchers and many financial institutions. These models are dedicated to measurement the default probability of credit portfolio. The evaluation of their default frequencies and the size of the credit portfolio are expressed as functions of macro-economic and micro-economic conditions as well as unobservable credit risk factors, which explained by other factors. We developed three sections to explain the different characteristics of the structural models of credit portfolio management. The purpose of all its models is to express the probability of default.

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Haw the Default Probability is Defined by the Credit Portfolio Models: A Comparative Theoretical Analysis Between the Structural Models?

Abdelkader Derbali ^α & Slaheddine Hallara ^σ

Abstract - This paper is elaborate of which the main is to present a theoretical analysis between the structural models. There are currently three types of models to consider the risk of credit: the structural models (The KMV Moody's model and the CreditMetrics model) also defined by the models of the value of the firm, reduced form models also defined by models with intensity (actuarial models) and the econometric models (The macro-factors models). The development of its three types of models is based on a theoretical basis developed by several researchers and many financial institutions. These models are dedicated to measurement the default probability of credit portfolio. The evaluation of their default frequencies and the size of the credit portfolio are expressed as functions of macro-economic and micro-economic conditions as well as unobservable credit risk factors, which explained by other factors. We developed three sections to explain the different characteristics of the structural models of credit portfolio management. The purpose of all its models is to express the probability of default.

Keywords : risk management; credit risk; default probability; structural models; kmv model; creditmetrics model.

I. INTRODUCTION

The problem of evaluation of the failure probability of any borrower was the center of the bankers as soon as they began to lend some money. The quantitative modeling of the credit risk for a debtor is rather recent in fact. Besides, the modeling of the credit risk associated with instruments of a portfolio of credit such as, the loans, the pledges, the guarantees and the by-products (who constitute a recent concept).

A certain number of models were developed, including at the same time the applications of property developed for the internal custom by the financial institutions, and the applications intended for the sale or for the distribution (Hickman and Koyluoglu, 1999).

The big financial institutions recognize his necessity, but there is a variety of approaches and rival

methods. There are three types of models of credit portfolio in the course of use at present (Crouhy et al., 2000).

- The structural models: there are two models of management of credit portfolio who are supplied in the literature: Moody's KMV model (Portfolio Model) and CreditMetrics model by JPMorgan.
- The Macro-factors model (Econometric model): The Credit Portfolio View model introduces in 1998 by Mckinsey.
- The actuarial models CSFP (Credit Suisse First Boston): this model (CreditRisk+) is developed in 1997.

The main idea for this study is to answer the question follows: *Haw the default probability is defined by the credit portfolio models?*

Then, the organization of this paper is as follows. In section 1, we present the main idea of the structural models and we define the forces and the weaknesses of each model. We provide the presentation of the KMV models and we define those forces and the weaknesses in section 2. The section 3 is considered to present the development of the Credit Metrics models and define those forces and the weaknesses. The final section is our conclusion.

II. WHAT ARE THE STRUCTURAL MODELS?

The structural models of management of credit portfolio were presented by Merton (1974) and then, developed by Leland (1994), Leland and Toft (1996), Anderson and Sundaresan (1996) and Jarrow (2011). The characteristics to define a structural model are given by two conditions:

- The process of management of the assets of the company has to be known on the market in which this one operates.
- The structure of the liabilities of the company has to be known by all the actors operating on the market of this one.

In the practice, to examine the models of management of credit portfolio, it is necessary to use parameters estimated implicitly because the values of

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the assets of the company are not observable. Nevertheless, the majority of the empirical evidence does not retain the structural models. The implicit prices obtained from the structural models does not seem to match the structure of maturity of the efficiencies on the assets of the company (*Eom et al., 2004; Ericsson and Reneby, 2005; Jarrow et al., 2003; Schaefer and Strebulaev, 2008; Li and Wong, 2008; Jarrow, 2011*) and to allow the forecasts of defect of the borrowers (*Patel and Pereira, 2007; Bharath and Shumway, 2008*).

The analysis of the model of *Merton (1974)* shows that this one supposes that the value of the firm follows a process of distribution and the defect occurs when the value of the firm falls below the nominal value of the debt on the date of maturity of this one. In this respect, this model serves to determine a threshold of defect.

The development of Merton's model is made by adding the other variables such as; the interest rate (*Longstaff and Schwartz, 1995*), the optimal permanent capital (*Leland and Toft, 1996*), the variable time of the

threshold of default (*Collin-Dufresne and Goldstein, 2001*), the unfinished accounting information (*Duffie and Lando, 2001*) and the risk of the events of defect (*Driessen, 2005*).

The structural models are based on the theory of the options and the structure of the capital of the company (*Hamisultane, 2008*). In this aligned, the bankruptcy of a company took place when the value of assets is situated below the value of its debt. The structural models or the models of the value of the firm are based on the approach of *Merton (1974)* which supposes that the failure of a company appears in case the market value of its assets is lower than a certain threshold of its debts.

Generally, the models of credit portfolio management resting on the approach of Merton are the model KMV (*Kealhofer, McQuown and Vasicek*) of Moody and the CreditMetrics model of JPMorgan (*1997*). The distinction between both structural models was described in the table below.

Table 1 : The comparison between the KMV model and the CreditMetrics model

The KMV approach	The CreditMetrics approach
▪ The conduct of the value of the asset. ▪ Companies are decomposed into systematic components and that no-systematic. ▪ The systematic risk is based on the industry and the country of debtor. ▪ The correlation of defect ensues from the correlation of assets.	▪ The indication of own capital. ▪ Companies are decomposed systematic components and that no-systematic. ▪ The systematic risk is based on the industry and the country of debtor and can be sensitive to the size of the asset. ▪ The correlation of the defect ensues from the correlation of the efficiencies on own capital.

Source : Smithson (2003)

The structural models are also called models of the asset volatility. The Structural aspect of the models comes because there is a historical story behind by default that is something manages to start by default. The structural models are rooted in the knowledge of Merton. In Merton's model, the correlation of defect has to be a function of correlation of assets. The estimation of a structural model requires the implementation of the market value of the assets of the company and its volatility.

In the practice, the value of assets and their volatility are not observable for the most part of companies. The structural models lean strongly on the existence of assets quoted on the stock exchange so that we can estimate the necessary parameters.

III. THE KMV MODEL

The KMV model of credit portfolio management was elaborated for the first time in 1993. This model allowed the development of several models of quantification of the credit risk: Credit Monitor, Credit

Edge and Private Firm Model for the individual credit risk and Portfolio Manager for the credit risk of a portfolio.

The model KMV rests bases on the notion of default distance which is calculated by basing itself on the barrier which engages the defect. As soon as, the distance in the defect is calculated, it transformed into the probability of failure (Expected Default Frequency: EDF).

The KMV model which was developed by the Moody's-KMV company is based on the theory of the prices of Merton options. It is about an abstract frame used to estimate the default probability of a company. The KMV model supposes that the company is in situation of defect when the value of its asset is less than the value of its debts. The Figure 1 explains the relation between the estimated own capital and the value of the asset. According to Merton's basic idea, in the KMV model the value of the own capital of the company is considered as being an option to buy. So, the value of the asset is considered as being the underlying asset and the debt represents the price of exercise (*Chen et al., 2010*).

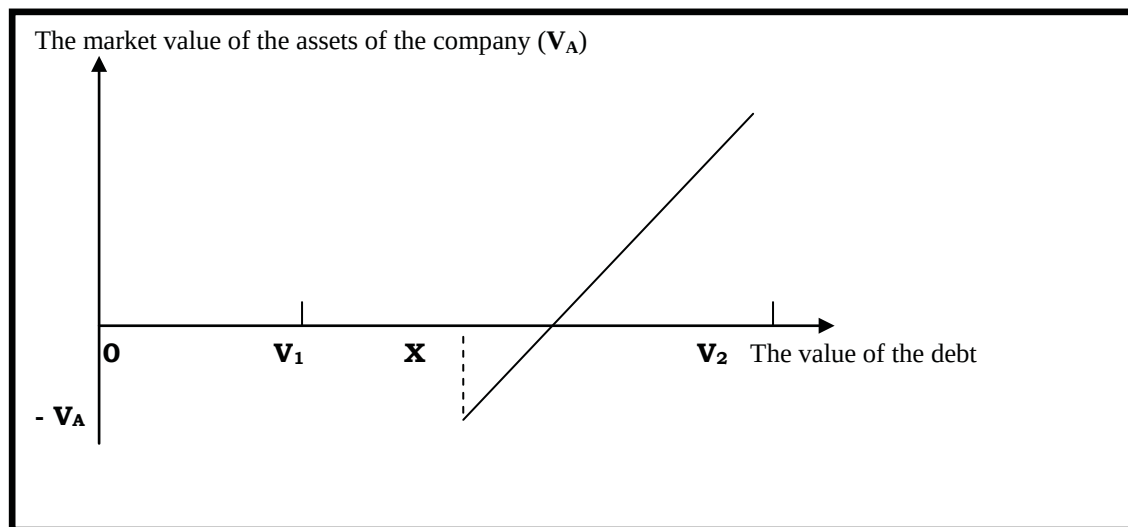


Figure 1 : The relation between the market value of the assets of the company and the value of the debt
(Merton, 1974)

In the Figure 1, V_A indicates the initial investment of the shareholders of the company; X indicates the point of default which corresponds to the sum of the long-term debt and half of the current liabilities. When the value of assets (V_A) is superior to the debt (X), the shareholders will choose to gain profits staying after payment of the debts ($V_A - X$) and these will be chosen by default, what is shaped with a net value raised in the Figure 1. In this case, the investor executes the option to buy.

So, if the value of assets is lower than the debt ($V_A < X$), the shareholders will choose by default the transfer of the active total for the benefit of the creditors, what is coherent with a constant value of own capital indicated in the Figure 1, and it means that the option to buy is not executed (Caouandte et al., 1998; Kealhofer and Bohn, 2001; Saunders and Allen, 2002; Bohn and Crosbie, 2003).

Generally speaking, the shareholders receive $\text{Max}(V_A - X, 0)$ in the date of maturity T . According to

Merton's model, the evolution of the market value of the assets of the company follows a process of geometrical distribution of the following shape:

$$\frac{dV_A}{V_A} = \mu dt + \sigma_A dW_t$$

Where W_t the process of Wiener Standard is, μ is the average of the efficiency of assets and σ_A is the standard deviation of the efficiency on assets. The market value of the company is given by basing itself on the purchase price of a European option to buy supplies by Black and Scholes (1973).

$$V_E = V_A N(d_1) - e^{-rT} X N(d_2)$$

Where $N(\cdot)$ Indicate the function of distribution of the normal law with (Huang and Yu, 2010):

$$\left\{ \begin{array}{l} d_1 = \frac{\left(\ln\left(\frac{V_A}{X}\right) + \left(r + \frac{1}{2}\sigma_A^2\right)T \right)}{\sigma_A \sqrt{T}} = \frac{1}{\sigma_A \sqrt{T}} \left(\ln\left(\frac{V_A}{X}\right) + \left(r + \frac{1}{2}\sigma_A^2\right)T \right) \\ d_2 = \frac{\left(\ln\left(\frac{V_A}{X}\right) + \left(r - \frac{1}{2}\sigma_A^2\right)T \right)}{\sigma_A \sqrt{T}} = d_1 - \sigma_A \sqrt{T} \end{array} \right.$$

In the KMV model, there is a hypothesis which rests on the structure of the capital of the company. So, this capital has to consist only by actions, current liabilities and in the long term and convertible prices. Really, the value of the company V_A and the volatility of assets σ_A are not observable (Hull, 1997; Chen et al.

2010). We are going to deduct these two values by using the values of the options V_E .

So land us note that: $V_E = f(V_A, \sigma_A, X, c, r)$

$$\sigma_E = g(V_A, \sigma_A, X, c, r)$$

Where c is the coupon paid on the long-term debt, r is the interest rate without the risk and σ_E is the volatility of share prices.

By applying the Lemma of Itô to these two functions and by arranging the terms we obtain:

$$\sigma_E = \left(\frac{V_A}{V_E} \right) \frac{\partial V_E}{\partial V_A} \sigma_A$$

With: who is deducted from the equation which measures the value of the VE which is defined by the following expression:

$$V_E = V_A N(d_1) - e^{-rT} X N(d_2)$$

Thus:

$$\sigma_E = \left(\frac{V_A N(d_1)}{V_E} \right) \sigma_A$$

Further to this transformation, we obtain a system of equation to two unknowns V_A and σ_A :

$$\begin{cases} V_A N(d_1) - e^{-rT} X N(d_2) - V_E = 0 \\ \sigma_E V_E - V_A N(d_1) \sigma_A = 0 \end{cases}$$

$$P_{KMV} = \text{Prob}\{V_A(T) < X\} = N \left(- \frac{\ln \left(\frac{V_A}{X} \right) + \left(\mu - \frac{1}{2} \sigma_A^2 \right) T}{\sigma_A \sqrt{T}} \right) = N(-DD)$$

Then we can obtain the frequency planned by default (Expected Default Frequency: EDF) such as:

$$EDF = N(-DD)$$

However, the default probability does not correspond to the normal law. KMV Company tries to obtain the empirical value of the EDF rather than the theoretical value of the models (Zheng, 2005).

Fortunately, KMV Company possesses an enormous base of historical data concerning the default of the companies. By basing itself on these data KMV defined tables which associate with the various possible values of the distance of default (DD) on a temporal horizon considered a default probability definite and noticed empirically.

If the expressions of V_A and σ_A are determined, then we can arrive at the writing of the following formulation of the distance of defect (DD):

$$DD = \frac{\ln \left(\frac{V_A}{X} \right) + \left(\mu - \frac{1}{2} \sigma_A^2 \right)}{\sigma_A \sqrt{T}}$$

According to the KMV model the distance of defect is defined in the following way (Crosbie and Bohn, 2003):

$$DD = \frac{V_A - X}{\sigma_A V_A}$$

From the distance of defect, we can deduct the value of the default probability as follows:

So, to protect itself against the risk which results from potential losses bound to the evolutions of the portfolio, Kealhofer, McQuown and Vasicek (1993) based on the determination of a random size L relative to the losses of the portfolio which is defined in a general way and on a horizon H as follows:

$$L = \frac{V_H}{ND} - V_H$$

Where $\frac{V_H}{ND}$ indicates the value of the portfolio

H in the absence of the losses and V_H indicates the market value of the portfolio H . The development follows by KMV shows us that the distribution of L can be approached by an inverse normal distribution.

Table 2 : The forces and the weaknesses relative to the KMV model

The forces	The weaknesses
- The default probability is connected with the information of the market. - Contrary to CreditMetrics models and CreditRisk+ models the debtors are specific. We can distinguish them by basing itself on their default probability, on their own structure of capital and on their own	- A hypothesis which is not realistic because she supposes that the debt of the company consists by bonds with zero-coupon and shares. - KMV supposes that the price of assets follows one moment Geometric Brownian. This modeling by a continuous process excludes all the early defaults.

assets. ▪ The threshold of defect is determined in an empirical way.	▪ This method is difficult because it depend a several data which are in most of the time unobservable or with difficulty accessible. ▪ The interest rate is supposed constant.
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Source : Hamisultane (2008)

IV. THE CREDITMETRICS MODEL

CreditMetrics was thrown for the first time in 1997 by JP Morgan's bank. CreditMetrics is considered as being an evaluation tool, for a portfolio, its variance of the values provoked by the changes of the quality of credit of the transmitter of the bonds (the credit migration) and leaves the defect of the counterpart. Unlike the approaches developed by the other models of management of a portfolio of credit, the probability of default in CreditMetrics is given by rating agencies for the big companies and by methods of scoring and mapping for small and medium-sized firms (Paleologo et al., 2010).

CreditMetrics belongs to the structural models since it rests on the model of Merton (1974) for the definition of the thresholds of the migration of credit (Jarrow, 2011). According to Hamisultane (2008), CreditMetrics makes it possible to calculate CreditVaR

of a portfolio. The methodology of this model is based on the probability of moving of a quality of credit to the other in a given horizon of time (analysis of the migration of credit). The calculation of CreditVaR by CreditMetrics rests on the four stages following (Crouhy et al., 2000; Hamisultane, 2008):

- Determination of the risk isolated from each credit of the portfolio.
- The construction of the matrix of the probabilities of transition from a notation to another.
- The valuation of the assets of the portfolio according to the scenarios of transition from a notation to the other one.
- The calculation of CreditVaR.

The evaluation of a portfolio Value-at-Risk due to the credit (CreditVaR) by CreditMetrics is given the following Figure 2 (Crouhy et al., 2000):

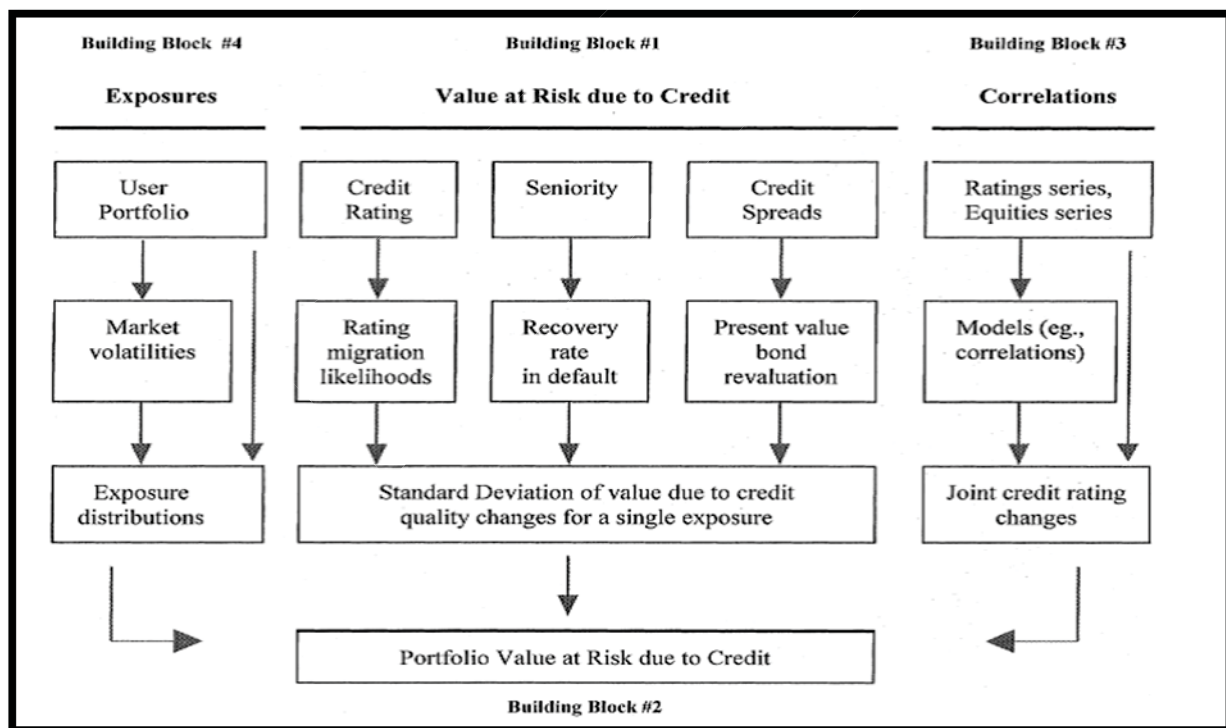


Figure 2 : The evaluation of a portfolio

In the model CreditMetrics, there are three categories of estimation to be used according to the nature of the composition of the portfolio. We are going to try, in what follows, to present the various principles of the model according to the composition of the portfolio.

a) The portfolio in an obligation

According to Hamisultane (2008), the system of rating used by CreditMetrics is the one rating agency. So, the broadcasting issuers of debt securities are noted according to a ladder of seven categories going

from AAA to CCC according to the financial solidity of every company (*Crouhy et al., 2000*). The notation AAA is tuned to the healthy companies financially whereas those who are characterized by a bad financial situation are noted by CCC.

The notations offered by the agencies of rating are regularly published. These notations present information relative to the broadcasting issuers of debt securities. The agencies of rating include these notations in indicating tables, either the rate of historic

default of broadcasting issuers according to their notation on a horizon of well determined time, or the evolutions of these notations in the time. These tables recapitulating the notations tuned to the broadcasting issuers of debt securities are defined by "the matrices of transition".

The matrices of annual transition summarize all the changes of notation, on a horizon of time of one year, of a sand of broadcasting issuers is presented as follows:

Table 3 : Transition matrix: Probabilities of credit rating migrating from one rating quality to another, within 1 year

Rating	AAA	AA	A	BBB	BB	B	CCC	Default
AAA	90.81%	8.33%	0.68%	0.06%	0.12%	0.00%	0.00%	0.00%
AA	0.70%	90.65%	7.79%	0.64%	0.06%	0.14%	0.02%	0.00%
A	0.09%	2.27%	91.05%	5.52%	0.74%	0.26%	0.01%	0.06%
BBB	0.02%	0.33%	5.95%	86.93%	5.30%	1.17%	0.12%	0.18%
BB	0.02%	0.14%	0.67%	7.73%	80.53%	8.84%	1.00%	1.06%
B	0.00%	0.11%	0.24%	0.43%	6.48%	83.46%	4.08%	5.20%
CCC	0.22%	0.00%	0.22%	1.30%	2.38%	5.00%	64.85%	19.79%
Default	0.00%	0.00%	0.00%	0.00%	0.00%	0.00%	0.00%	100%

Source : Standard & Poor's CreditWeek (1996)

According to *Grundke (2009)*, this table must be carefully analyzed. So, by taking as an example the line corresponding to the BBB rating presented in the table above, we can deduct the probability of default as follows:

Table 4 : The potential rating relative to the BBB rating

Initial rating	Potential rating in a one year	Probability
BBB	AAA	0.02%
	AA	0.33%
	A	5.95%
	BBB	86.93%
	BB	5.30%
	B	1.17%
	CCC	0.12%
	D	0.18%
	Total	100.00%

Source : *Grundke (2009)*

After a period of one year, and settling on the asset of initial notation BBB, we can deduct that the probability that this active rest BBB after a period of one year is 86,93 %, that to become AAA is 0,02 % and that to be lacking is 0,18 %.

The use of this model is based on three main hypotheses (*JP Morgan and Co. Inc, 1997; Glasserman*

and *Li, 2005; Hamisultane, 2008; Grundke, 2009; Figlewski et al., 2012*):

- The absence of multiple transitions: for a horizon of time given the number of transitions is in most of a single transition.
- The stability of the matrix of transition in time: for every class of notation, two companies in different sectors or in different countries have the same probability to migrate from a notation to the other one.
- The matrix of transition is of type Markov: for period given the probability to migrate of a class of notation in another class is independent from what took place for the last periods. These hypotheses are emitted for the simplification of the calculations of the matrix of transition for the posterior periods.

CreditMetrics determines the current value of the bond by using the curve of the rates with zero coupons to proceed with the calculations of CreditVaR. In that case, the transmitter of debt securities is not in situation of bankruptcy. By continuing in the same context of analysis, that is the use of the notation BBB as the example, we can use the table of the Forward rates following:

Table 5 : One-year forward zero-curves for each credit rating (%)

Category	Year 1	Year 2	Year 3	Year 4
AAA	3.60	4.17	4.73	5.12
AA	3.65	4.22	4.78	5.17
A	3.72	4.32	4.93	5.32
BBB	4.10	4.67	5.25	5.63

<i>BB</i>	5.55	6.02	6.78	7.27
<i>B</i>	6.05	7.02	8.03	8.52
<i>CCC</i>	15.5	15.02	14.03	13.52

Source : CreditMetrics, JP Morgan

We suppose in our case which a noted transmitter BBB has emitted a Bond for 100 Euro over 4

years with a rate without annual risk of 6 %. The current value of the bond is given by the equation below:

$$V = 6 + \frac{6}{(1+4.1\%)} + \frac{6}{(1+4.67\%)^2} + \frac{6}{(1+5.25\%)^3} + \frac{106}{(1+5.63\%)^4} = 107.55$$

By basing itself on the formula above, being able to us determine the various possible values of fire of type BBB according to his possible migrations towards other notations (*Crouhy et al., 2000; Hamisultane, 2008*). The possible values of a bond rated BBB according to the possible migrations are presented in the table 5.

In case the company had a bankruptcy, the value of the bond is determined by using the average

recovery ratio calculated by CreditMetrics on historical data (*Carty and Lieberman, 1996; Gordy, 1998*).

Further to the representative table of the various values of BBB according to the possible migrations, we can subtract the distribution of the variations of the price of the obligation in the following table:

Table 6 : Distribution of the bond values, and changes in value of a BBB bond, in 1 year

Rating	Probability: p (%)	Price of the obligation(bond) V (\$)	Difference with regard to V: ΔV	Difference with regard to the average μ	μ ² * p (%)
<i>AAA</i>	0.02	109.37	1.82	2.28	0.0010
<i>AA</i>	0.33	109.19	1.64	2.10	0.0146
<i>A</i>	5.95	108.66	1.11	1.57	0.1474
<i>BBB</i>	86.93	107.55	0	0.46	0.1853
<i>BB</i>	5.30	102.02	-5.53	-5.06	1.3592
<i>B</i>	1.17	98.10	-9.45	-8.99	0.9446
<i>C</i>	0.12	83.64	-23.91	-23.45	0.6598
<i>Default</i>	0.18	51.13	-56.42	-55.96	5.6358
	Average =	107.09 (\$)		Variance = Standard deviation =	8.9477 2.99 (\$)

Source : CreditMetrics, JP Morgan

The analysis of this table shows that CreditVaR in 1 % (at a level of 99 % confidence) is equal to the last value of the variation of the value of the bond which corresponds to the notation CCC. Thus, CreditVaR is equal to -23.91.

b) The portfolio in two obligations

In the case of a portfolio consisted of two bands, the analysis is based on the level of correlation of the migrations. In fact, in a portfolio consisted of several assets the migrations of the various credits are correlated. CreditMetrics tries to estimate these correlations. As long, as there are no good data to be used. In that case, CreditMetrics used the correlations between the values of the assets of the broadcasting issuers of the credits which are approached by the correlations between the equity prices of these broadcasting issuers to calculate the correlations between the migrations of the credits (*Treacy and Carey, 2000; Altman and Rijken, 2004; Gordy and Howells, 2006; Xing et al., 2012*).

According to *Iscoe et al. (1999)*, to be able to divert the correlations of the migrations of the credits of the correlations of the values of assets, it is necessary to have a model linking the quality of a credit to the value of assets. The solution proposed by CreditMetrics is to use an extension of the model of *Merton (1974)* which incorporates the migrations of the credits. In this aligned, we suggest taking into account the probability of migration of a bond rated initially by BB. These probabilities are given by the following table:

Table 7 : Transition matrix based on actual rating changes

Rating	Probability: p (%)	Price of the obligation(bond) V (\$)	Difference with regard to V: ΔV	Difference with regard to the average μ	μ ² * p (%)
AAA	0.02	109.37	1.82	2.28	0.0010
AA	0.33	109.19	1.64	2.10	0.0146
A	5.95	108.66	1.11	1.57	0.1474
BBB	86.93	107.55	0	0.46	0.1853
BB	5.30	102.02	-5.53	-5.06	1.3592
B	1.17	98.10	-9.45	-8.99	0.9446
C	0.12	83.64	-23.91	-23.45	0.6598
Default	0.18	51.13	-56.42	-55.96	5.6358
	Average =	107.09 (\$)		Variance = Standard deviation =	8.9477 2.99 (\$)

Source : CreditMetrics, JP Morgan

Year 2013

28

By basing itself on the model of Merton (1974), we can suppose that the efficiency on a bond modeled as follows:

$$r = \mu + \sigma \varepsilon$$

With: ε a term of error is such as $\varepsilon \sim N(0,1)$, μ is the average efficiency on the bond and σ is the standard deviation of the efficiencies of this bond. Then, the default probability of an issuer of the bond is given by the following expression:

$$Pr\{default\} = Pr\{r < Z_{Def}\} = Pr\{\mu + \sigma \varepsilon < Z_{Def}\}$$

Thus,

$$Pr\{default\} = Pr\{r < Z_{Def}\} = Pr\{\sigma \varepsilon < Z_{Def}\}$$

If $\mu=0$

$$Pr\{default\} = \left\{ \varepsilon < \frac{Z_{Def}}{\sigma} \right\} = \Phi\left(\frac{Z_{Def}}{\sigma}\right)$$

Where Φ indicates the cumulative function of the normal law.

By using the table above, we can establish the table according to who summarizes the distribution of the probability of migration affected in conformance with BB rating:

Table 8 : The distribution of the probability of migration of BB rating

Rating	Probability from the transition matrix (%)	Probability according to the asset value model
AAA	0.03	$1 - \Phi(Z_{AA}/\sigma)$
AA	0.14	$\Phi(Z_{AA}/\sigma) - \Phi(Z_A/\sigma)$
A	0.67	$\Phi(Z_A/\sigma) - \Phi(Z_{BBB}/\sigma)$
BBB	7.73	$\Phi(Z_{BBB}/\sigma) - \Phi(Z_{BB}/\sigma)$
BB	80.53	$\Phi(Z_{BB}/\sigma) - \Phi(Z_B/\sigma)$
B	8.84	$\Phi(Z_B/\sigma) - \Phi(Z_{CCC}/\sigma)$
CCC	1.00	$\Phi(Z_{CCC}/\sigma) - \Phi(Z_{Def}/\sigma)$
Default	1.06	$\Phi(Z_{Def}/\sigma)$

Source : Crouhy and al. (2000)

With, $1 - \Phi\left(\frac{Z_{AA}}{\sigma}\right)$ represent the probability so

that the bond of BB rating can pass in the notation AAA and Z_{AA} indicates the threshold from which BB passes to A_{AA} .

The transformation graphic of the data above is presented as follow:

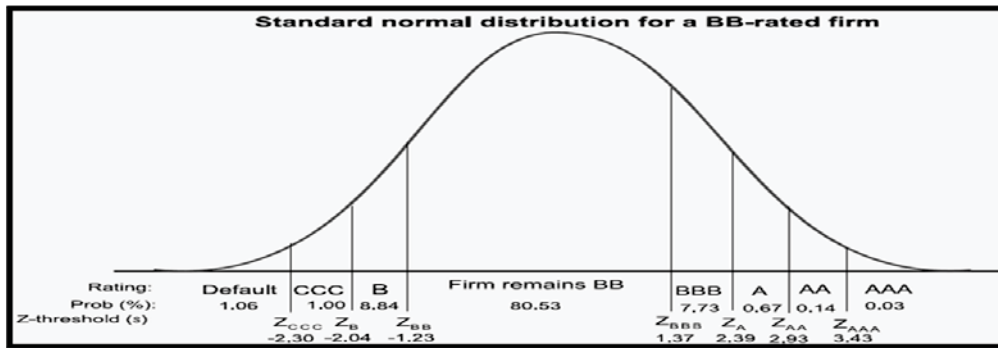


Figure 3 : Generalization of the Merton model to include rating changes (Crouhy and al., 2000)

Thus:

$$Z_{def}\sigma = \Phi^{-1}(1.06\%) = -2.30$$

The values of the other thresholds are calculated according to whom corresponds itself aside type of the normal distribution of the random on the assets of the notation BB (Gupton et al., 1997; Crouhy et al., 2000; Nickell et al., 2000; Bangia et al., 2002;

Albanese and Chen, 2003; Albanese et al., 2003; Rosch, 2005; Feng et al., 2008).

We suppose now, that a second issuer presents a rating A where the random on assets follow a normal distribution with a parameter σ^A . In that case, the values of thresholds relative for two bands who rated BB and A are presented as follows:

Table 9 : Transition probabilities and credit quality thresholds for BB and A rated obligors

Rating in 1 year	Rated-A obligor		Rated-BB obligor	
	Probabilities (%)	Thresholds: $Z_{(\sigma)}$	Probabilities (%)	Thresholds : $Z_{(\sigma)}$
AAA	0.09	3.12	0.03	3.43
AA	2.27	1.98	0.14	2.93
A	91.05	-1.51	0.67	2.39
BBB	5.52	-2.30	7.73	1.37
BB	0.74	-2.72	80.53	-1.23
B	0.26	-3.19	8.84	-2.04
CCC	0.01	-3.24	1.00	-2.30
Default	0.06		1.06	

Source: Crouhy and al. (2000)

By taking into account the table above, we can calculate the probability of migration joined in the following way:

$$P(Z_{BB} < r < Z_{BBB}, Z_A < r' < Z_{AA}) = \int_{Z_{BB}}^{Z_{BBB}} \int_{Z_A}^{Z_{AA}} f(r, r', \sigma, \sigma') dr dr'$$

With r and r' indicate respectively the random on the assets who are rated by BB and A and $f(r, r', \sigma, \sigma')$ represent the joint density function by the Gaussian distribution which depends on the coefficient of correlation ρ .

The joint density function of the Gaussian distribution of two variables X and Y is presented by the form below:

$$f(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)}\left(\frac{x^2}{\sigma_x^2} + \frac{y^2}{\sigma_y^2} - \frac{2\rho xy}{\sigma_x\sigma_y}\right)\right)$$

According to Hamisultane (2008), for $\rho = 20\%$ the matrix of joint transition which considers the correlation banding both entities BB and A is the following one:

Table 10 : Joint rating probabilities (%) for BB and A rated obligors when correlation banding asset random is 20%

Rating of first company (BB)	Rating of second company (A)								
	AAA	AA	A	BBB	BB	B	CCC	Default	Total
AAA	0.00	0.00	0.03	0.00	0.00	0.00	0.00	0.00	0.03
AA	0.00	0.01	0.13	0.00	0.00	0.00	0.00	0.00	0.14
A	0.00	0.04	0.61	0.01	0.00	0.00	0.00	0.00	0.67
BBB	0.02	0.35	7.10	0.20	0.02	0.01	0.00	0.00	7.69
BB	0.07	1.79	73.65	4.24	0.56	0.18	0.01	0.04	80.53
B	0.00	0.08	7.80	0.79	0.13	0.05	0.00	0.01	8.87
CCC	0.00	0.01	0.85	0.11	0.02	0.01	0.00	0.00	1.00
Default	0.00	0.01	0.90	0.13	0.02	0.01	0.00	0.00	1.07
Total	0.09	2.29	91.06	5.48	0.75	0.26	0.01	0.06	100

Source : CreditMetrics, JP Morgan (Lucas, 1995)

The last column of the table and the last line of this one represent the marginal probability for the entities BB and A which are equal to the sum of the joint probability by line or by the column. According to Crouhy and al. (2000) these marginal probabilities correspond to the probability of migration of BB and of A taken individually. The variation of the portfolio of both bands is calculated for each of the joint probability (Brady and Bos, 2002; Brady and al., 2003).

c) The portfolio in several obligations

In case the portfolio consists further more than 2 bands calculates its joint probability will more be complicated. So, model CreditMetrics propose the use of the simulations of Monte Carlo and the decomposition of Cholesky to generate trajectories correlated to the bond and build the distribution of the values of the portfolio on certain horizon of time (Gouriéroux and Monfort, 1995; Fishmen, 1997; Crouhy et al., 2000; Hamilton et al., 2002).

$$cov(r_i, r_j) = \hat{\alpha}_{iX} \hat{\alpha}_{jX} V(I_X) + \hat{\alpha}_{iY} \hat{\alpha}_{jY} V(I_Y) + (\hat{\alpha}_{iX} \hat{\alpha}_{jY} + \hat{\alpha}_{iY} \hat{\alpha}_{jX}) cov(I_X, I_Y)$$

And

$$V(r_i) = \hat{\alpha}_{iX}^2 V(I_X) + \hat{\alpha}_{iY}^2 V(I_Y) + V(v_i^2) 2(\hat{\alpha}_{iX} \hat{\alpha}_{iY}) cov(I_X, I_Y)$$

By using these two formulae, we can obtain the matrix of the variances-covariance's Σ .

Stage 3 : The decomposition of Cholesky of the matrix of the variances of the variances-covariance's Σ in the following way (Hamisultane, 2008):

$$\Sigma = AA^T$$

With A represent the lower triangular matrix and AT transposed by the matrix A.

Stage 4: The simulation of variables v $Z_{i,t} \sim N(0,1)$. In fact, the existence of the bond to be feigned allows the existence of $f Z_{i,t}$.

According to Hamisultane (2008) and Feng et al. (2008), to generate trajectories correlated to the variables which follow a normal distribution $N(\mu, \Sigma)$. The determination of these trajectories requires the respect for the following five stages:

Stage 1: The regression of the random $r(t)$ of the band on the sectorial indications. For example, in the case of three bands and two sectorial indications.

$$r_{1,t} = \alpha_1 + \alpha_{1,X} I_{X,t} + \alpha_{1,Y} I_{Y,t} + v_{1,t}$$

$$r_{2,t} = \alpha_2 + \alpha_{2,X} I_{X,t} + \alpha_{2,Y} I_{Y,t} + v_{2,t}$$

$$r_{3,t} = \alpha_3 + \alpha_{3,X} I_{X,t} + \alpha_{3,Y} I_{Y,t} + v_{3,t}$$

To pass in the second stage it is necessary to estimate the various parameters of three models.

Stage 2: The calculation of the variances and the covariance's banding 2 bands i and j:

Stage 5: The simulation of the values of the correlated variables by basing itself on a process of geometrical distribution:

$$\frac{dV}{V} = \mu dt + A \sqrt{dt} Z$$

Thus:

$$\frac{dV}{V} = \begin{pmatrix} dV_t^1/V_t^1 \\ dV_t^2/V_t^2 \\ \vdots \\ dV_t^i/V_t^i \\ \vdots \\ dV_t^n/V_t^n \end{pmatrix} \approx \begin{pmatrix} \ln V_t^1 - \ln V_{t-1}^1 \\ \ln V_t^2 - \ln V_{t-1}^2 \\ \vdots \\ \ln V_t^i - \ln V_{t-1}^i \\ \vdots \\ \ln V_t^n - \ln V_{t-1}^n \end{pmatrix}$$

$$\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_i \\ \vdots \\ \mu_n \end{pmatrix}$$

$$dt = \Delta t$$

$$A = \begin{pmatrix} \beta_{11} & 0 & \dots & 0 \\ \beta_{21} & \ddots & & 0 \\ \vdots & & \ddots & \vdots \\ \beta_{i1} & & \beta & 0 \\ \vdots & & \beta & \vdots \\ \beta_{n1} & \beta_{n2} & \dots & \beta_{nn} \end{pmatrix}$$

$$\sqrt{dt} Z \approx \sqrt{\Delta t} \begin{pmatrix} Z_{1,t} \\ Z_{2,t} \\ \vdots \\ Z_{i,t} \\ \vdots \\ Z_{n,t} \end{pmatrix}$$

According to Crouhy et al. (2000), Nickell et al. (2000) and Bangia et al. (2002), the forces and the weaknesses of this model are presented in the following table:

Table 11 : The forces and the weaknesses relative to the CreditMetrics model

The forces	The weaknesses
In the model CreditMetrics, both aspects of the credit risk are taken into account.	- The rating according to companies must be correct. - The interest rates are supposed constant. - The existence of a relation between the economic situation and the probability of defect. In that case, every economic cycle has to have matrices of transition appropriate for him. - The variability of the actions of a company can be used to deduct the variability of the price of the assets of the company.

Source: Crouhy and al. (2000), Nickell and al. (2000) and Bangia and al. (2002)

V. CONCLUSION

In this paper we developed a comparative theoretical approach's concerning the model of management of credit portfolio. Then, we studied the four mains models of credit portfolio management. In the financial literature those models are grouped by three types of credit portfolio models (Crouhy et al., 2000). The structural models: there are two models of management of credit portfolio who are supplied in the literature: Moody's KMV model (Portfolio Model) and CreditMetrics model by JPMorgan.

The KMV model and Credit Portfolio View base their approach on the same empirical observation that default and migration probabilities vary over time. The KMV model adopts a microeconomic approach which relates the probability of default of any obligor, to the market value of its assets. The Credit Portfolio View model proposes a methodology which links macroeconomics factors to default and migration probabilities. The calibration of this model necessitates reliable default data for each country, and possibly for each industry sector within each country.

Structural models are based on option theory and capital structure the company. On econometric

models, they link the probability fault of the company to the state of the economy. The probability of failure depends in these models of macroeconomic factors such as unemployment, the rate of increase GDP, the interest rate long-term. Moreover, in the CreditRisk+ models, the probability of default varies over time.

REFERENCES RÉFÉRENCES REFERENCIAS

1. Ali, A. and Daly, K. (2010). Macroeconomic determinants of credit risk: Recent evidence from a cross country study. *International Review of Financial Analysis*, (19):165–171.
2. Allen, L. and Saunders, S. (3003). A survey of cyclical effects in credit risk measurement models. *BIS Working Paper*, No. 126, New York University.
3. Bensoussan, A., Crouhy, M. and Galai, D. (1995). Stochastic equity volatility related to the leverage effect II: Valuation of European equity options and warrants. *Applied Mathematical Finance*, Vol. 2, pp.43-59.
4. Berry, M., Burmeister, E. and McElroy, M. (1998). Sorting our risks using known APT factors. *Financial Analysts Journal*, 44 (2):29-42.

5. Credit Suisse Financial Products (1997). CreditRisk+: A Credit Risk Management Framework.
6. Crouhy, M., Galai, D. and Mark, R. (2000). A comparative analysis of current credit risk models. *Journal of Banking & Finance*, (24):59-117.
7. Figlewski, S., Frydman, H. and Liang, W. (2012). Modeling the effect of macroeconomic factors on corporate default and credit rating transitions. *International Review of Economics and Finance*, (21):87-105.
8. Grundke, P. (2005). Risk Measurement with Integrated Market and Credit Portfolio Models. *Journal of Risk*, 7 (3):63-94.
9. Grundke, P. (2009). Importance sampling for integrated market and credit portfolio models. *European Journal of Operational Research*, (194):206-226.
10. Gupton, G.M., Finger, C.C. and Bhatia, M. (1997). *CreditMetricsTM – Technical Document*, Morgan Guaranty Trust Company.
11. Hamisultane, H. (2008). *Modèles de gestion du risque de crédit. Investment System R&D, Document n°1*.
12. Huang, S.J. and Yu, J. (2010). Bayesian analysis of structural credit risk models with microstructure noises. *Journal of Economic Dynamics & Control*, (34):2259-2272.
13. Jarrow, R. and Turnbull, S. (1995). Pricing derivatives on financial securities subject to credit risk. *The Journal of Finance*, (50):53-85.
14. Jarrow, R.A., Lando, D. and Yu. F. (2001). Default risk and diversification: theory and applications. *Mathematical Finance*, (15):1-26.
15. Jarrow, R.A. (2011). Credit market equilibrium theory and evidence: Revisiting the structural versus reduced form credit risk model debate. *Finance Research Letters*, (8):2-7.
16. Lee, W.C. (2011). Redefinition of the KMV model's optimal default point based on genetic-algorithms–Evidence from Taiwan. *Expert Systems with Applications*, (38):10107-10113.
17. Liao, H.H., Chen, T.K. and Lu, C.W. (2009). Bank credit risk and structural credit models: Agency and information asymmetry perspectives. *Journal of Banking & Finance*, (33):1520-1530.
18. Merton, R. (1974). On the pricing of corporate debts: the risk structure of interest rates. *Journal of Finance*, (29):449-470.
19. Musto, D.K. and Souleles, N.S. (2006). A portfolio view of consumer credit. *Journal of Monetary Economics*, (53):59-84.
20. Tarashev, N. (2010). Measuring portfolio credit risk correctly: Why parameter uncertainty matters. *Journal of Banking & Finance*, (34):2065-2076.
21. Vetendorpe, A., Ho, N.D., Vetuffel, S. and Dooren, P.V. (2008). On The Parameterization of the CreditRisk+ Model for Estimating Credit Portfolio Risk. *Insurance: Mathematics and Economics*, 42(2):736-745.
22. Xiaohong, C., Xiaoding, W. and Desheng, W.D. (2010). Credit risk measurement and early warning of SMEs: An empirical study of listed SMEs in China' *Decision Support Systems*, (49):301-310.
23. Zhang, Q. and Wu, M. (2011). Credit Risk Migration Based on Jarrow-Turnbull Model. *Systems Engineering Procedia*, (2):49-59.